## Research Design - -Topic 10 Multiple Regression and Multiple Correlation © 2010 R. C. Gardner, Ph.D.

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Multiple Correlation was introduced by Yule (1897) as an extension of bivariate regression to assess linear relations involving a number of independent variables. The intent was to improve prediction over the bivariate case.

Since then, there have been many applications, including:

1. Establishment of a prediction equation
2. Selection of a subset of "predictors"
3. Analysis of variance
4. Curve filting
5. Assessing mediation
6. Assessing moderation
7. Path analysis

## Applications and Implications

It is often said that multiple correlation can be used to identify good predictors. This is not the case. Multiple correlation does not identify predictors of a criterion. It identifies variables that add to prediction. There is a difference. Note that:

The Pearson product moment correlation between a variable and the criterion can be considered a measure of prediction. The correlation coefficient is the regression coefficient in standard score form.

The regression coefficient in multiple regression is a measure of the extent to which a variable adds to the prediction of a criterion, given the other variables in the equation. It is not a correlation coefficient.

Multiple regression is an equation linking a criterion variable (X) to a set of other variables. For example, one might wish to predict grades in a subject (the criterion) with a number of other variables such as GRE-Verbal, GRE-Quantitative, and Height.

The general form of the regression equation in raw score form is:

$$
X^{\prime}=b_{0}+b_{1} V_{1}+b_{2} V_{2}+\ldots+b_{k} V_{k}
$$

In standard score form, the equation is:

$$
Z_{X}^{\prime}=\beta_{1} Z_{1}+\beta_{2} Z_{2}+\ldots+\beta_{k} Z_{k}
$$

Multiple Correlation is the Pearson product moment correlation of the obtained and predicted values of $X$.

$$
R=\frac{\sum(X-\bar{X})\left(X^{\prime}-\bar{X}\right)}{n S_{X} S_{X^{\prime}}}
$$

And with a bit of algebra: $\quad R=\sqrt{\beta_{1} r_{1 x}+\beta_{2} r_{2 x}+\ldots+\beta_{k} r_{k x}}$
(i.e., the multiple correlation is equal to the square root of the sum of the product of the standardized regression coefficient for each predictor times its correlation with the criterion.)

$$
\text { And that: } \quad \beta_{k}=\frac{b_{k} S_{k}}{S_{X}}
$$

(i.e., the standardized regression coefficient is equal to the unstandardized regression coefficient times the standard deviation of the predictor divided by the standard deviation of the criterion.)

## Basic Equations:

$$
\begin{array}{ll}
\text { Raw score form } & X_{1}^{\prime}=b_{0}+b_{2} X_{2}+b_{3} X_{3}+\ldots+b_{k} X_{k} \\
\text { Standard score form } & Z_{1}^{\prime}=\beta_{2} Z_{2}+\beta_{3} Z_{3}+\ldots+\beta_{k} Z_{k}
\end{array}
$$

The square of the multiple correlation is equal to the variance of the predicted $Z$ scores such that:

$$
\begin{aligned}
R^{2} & =\frac{\sum Z_{1}^{2}}{N}=\frac{\sum\left(\beta_{2} Z_{2}+\beta_{3} Z_{3}+\beta_{4} Z_{4}\right)^{2}}{N} \\
& =\beta_{2}^{2}+\beta_{3}^{2}+\beta_{4}^{2}+2 \beta_{2} \beta_{3} r_{23}+2 \beta_{2} \beta_{4} r_{24}+2 \beta_{3} \beta_{4} r_{34}
\end{aligned}
$$

i.e., the square of the multiple correlation is equal to the sum of the squared standardized regression coefficients plus two times the product of the correlation between each pair of predictors times their regression coefficients.

Matrix equations. Matrix notation is often used with multiple regression and correlation. The following examples consider the use of 3 predictors.

The squared multiple correlation is written as:

$$
R_{1.234}^{2}=\beta_{2} r_{12}+\beta_{3} r_{13}+\beta_{4} r_{14}
$$

which can be expressed as the product of two vectors as:

$$
\begin{aligned}
& =\left[\begin{array}{lll}
r_{12} & r_{13} & r_{14}
\end{array}\right]\left[\begin{array}{l}
\beta_{2} \\
\beta_{3} \\
\beta_{4}
\end{array}\right]=R_{1 j} \beta_{j} \\
& \beta_{j}=R_{i j}^{-1}
\end{aligned} R_{j 1}\left[\begin{array}{l}
\end{array}\right.
$$

where $\beta_{\mathrm{j}}$ is defined as:
i.e., the product of the inverse of the matrix of correlations of the predictors with the vector of correlations of the criterion with the predictors. That is:

$$
R_{j j} R_{j j}^{-1}=I
$$

Thus, in matrix terms:

$$
R_{1.234}^{2}=R_{1 j} R_{j j}^{-1} R_{j 1}
$$

## The Case of Two Predictors

The regression equation describes a plot in three dimensional space as indicated on the next slide. The plot shows the model in raw score form based on the regression equation as
follows:
where

$$
X_{1}^{\prime}=C+b_{2} X_{2}+b_{3} X_{3}
$$

$$
b_{2}=\frac{S_{X_{1}} \beta_{2}}{S_{X_{2}}} \quad \text { and } \quad b_{3}=\frac{S_{X_{1}} \beta_{3}}{S_{X_{3}}}
$$

and

$$
C=\bar{X}_{1}-b_{2} S_{X_{2}}-b_{3} S_{X_{3}}
$$



In the diagram, X 2 and X 3 are shown to be orthogonal (i.e., independent of each other), but generally the predictors are correlated. Thus, to ensure independence, we calculate the regression coefficients on residualized variables. This involves the constructs of partial and semipartial correlation.

1. Partial Correlation - Plots in Standard Score Form

$$
r_{12.3}=\frac{\sum\left(Z_{1}-Z_{1}^{\prime}\right)\left(Z_{2}-Z_{2}^{\prime}\right)}{n S_{Z_{1}-Z_{1}} S_{Z_{2}-Z_{2}^{\prime}}}=\frac{r_{12}-r_{13} r_{23}}{\sqrt{1-r_{13}^{2}} \sqrt{1-r_{23}^{2}}}
$$



2. Semipartial (part) Correlation

$$
r_{1(2.3)}=\frac{\sum Z_{1}\left(Z_{2}-Z_{2}^{\prime}\right)}{n S_{Z_{2}-Z_{2}}}=\frac{r_{12}-r_{13} r_{23}}{\sqrt{1-r_{23}^{2}}}
$$

The regression equations in standard score form

$$
Z_{X_{1}}^{\prime}=\beta_{2} Z_{X_{2}}+\beta_{3} Z_{X_{3}}
$$

Where:

$$
\begin{aligned}
& \beta_{2}=\frac{r_{12}-r_{13} r_{23}}{1-r_{23}^{2}}=\frac{r_{1(2.3)}}{\sqrt{1-r_{23}^{2}}} \\
& \beta_{3}=\frac{r_{13}-r_{12} r_{23}}{1-r_{23}^{2}}=\frac{r_{1(3.2)}}{\sqrt{1-r_{23}^{2}}}
\end{aligned}
$$

Thus: Beta coefficients can be shown to equal the semipartial correlation of the criterion with a predictor divided by the standard error of estimate of that predictor in standard score form as predicted by the other predictor.

Note that:
$R_{1.23}=\sqrt{\beta_{2} r_{12}+\beta_{3} r_{13}}$ 11

Relation of Multiple Correlation to Relations Among Predictors
Other things being equal, it can be shown that the the multiple correlation decreases as the correlation between predictors increases. Consider the case where

$$
r_{12}=.6 \quad r_{13}=.5
$$

It can be shown that:

$$
r_{23}=r_{12} r_{13} \pm \sqrt{1-r_{12}^{2}-r_{13}^{2}+r_{12}^{2} r_{13}^{2}}
$$

Thus: $\quad r_{23}=.30 \pm \sqrt{.48} \quad \therefore-.393 \leq r_{23} \leq .993$
Thus, we can consider the values of $\beta_{2}, \beta_{3}$, and $R_{1.23}$ when $r_{23}$ varies from -. 30 to . 90 . Applying the formulae would produce the following answers

| $r_{23}$ | $\beta_{2}$ | $\beta_{3}$ | $R_{1.23}$ | $r_{Z 1\left(Z_{2}+Z_{3}\right)}$ |
| :---: | :---: | :---: | :---: | :---: |
| .30 | .824 | .747 | .932 | .930 |
| .20 | .729 | .646 | .872 | .870 |
| .00 | .600 | .500 | .781 | .781 |
| .20 | .521 | .396 | .715 | .710 |
| .40 | .476 | .310 | .664 | .657 |
| .60 | .469 | .219 | .625 | .615 |
| Correlation of Z1 |  |  |  |  |
| .80 | .556 | .056 | .601 | .580 |
| (see Topic 15) |  |  |  |  |
| .90 | .789 | .211 | .607 | .564 |

## The Case of Many Predictors

Multiple correlation and multiple regression can be conducted with any number of predictors though it is wise to keep the number manageable. It is generally recognized that the dependent variable should be continuous (and normal) but predictors can be both continuous and categorical (while distributional characteristics will influence the results). The following discussion will focus primarily on three predictors though the generalizations apply to any number of predictors.

Multiple regression can be performed by entering all of the predictors of interest in one step or by using a hierarchical method in which the researcher enters the predictors in some predetermined manner either one at a time or in groups. There are also a number of indirect methods where the computer enters the predictors in a manner determined by the data. Some of these are discussed below.

Relation of $\mathrm{R}^{2}$ to semipartial correlations

$$
\begin{aligned}
R_{1.234}^{2} & =\beta_{2} r_{12}+\beta_{3} r_{13}+\beta_{4} r_{14} \\
& =r_{12}^{2}+r_{1(3.2)}^{2}+r_{1(4.32)}^{2}
\end{aligned}
$$

It can be shown that:

$$
\text { Semipartial correlation }{ }^{2}=r_{1(4,23)}^{2}=\frac{t_{b_{4}}^{2}\left(1-R_{1.234}^{2}\right)}{N-3-1}
$$

and
Semipartial correlation ${ }^{2}=r_{1(3.2)}^{2}=\frac{t_{b_{3}}^{2}\left(1-R_{1.23}^{2}\right)}{N-2-1}$
or

$$
\text { Partial correlation }^{2}=r_{14.23}^{2}=\frac{r_{1(4.23)}^{2}}{1-R_{4.23}^{2}}
$$

Etc...

The test of significance of the multiple R is:

$$
F=\frac{R^{2} / p}{\left(1-R^{2}\right) /(N-p-1)} \quad \text { With df: v1 }=\mathrm{p}, \mathrm{v} 2=\mathrm{N}-\mathrm{p}-1
$$

The test of significance of the increase in the Multiple $R$ when adding variables to an existing regression equation is:

$$
F=\frac{\left(R_{2}^{2}-R_{1}^{2}\right) /\left(p_{2}-p_{1}\right)}{\left(1-R_{2}^{2}\right) /\left(N-p_{2}-1\right)} \quad \text { With df: } \begin{aligned}
v 1 & =p_{2}-p_{1} \\
v 2 & =N-p_{2}-1
\end{aligned}
$$

Where:

$$
\begin{gathered}
R_{2}^{2}=\mathrm{R}^{2} \text { with } \mathrm{p}_{2} \text { predictors } \quad R_{1}^{2}=\mathrm{R}^{2} \text { with } \mathrm{p}_{1} \text { predictors } \\
\qquad \begin{array}{c}
p_{2}>p_{1} \\
N=\text { total number of subjects }
\end{array}
\end{gathered}
$$

Tests of Significance of the Regression Coefficients

Test of significance of $b$ :

$$
t_{b}=\frac{b}{S E_{b}} @ d f=N-p-1
$$

It will be recalled that this is the square root of the $F$ for $R^{2}$ change when the predictor is added to the equation.

Test of significance of $\beta$

$$
t=\frac{\beta_{j}}{S E_{\beta j}}=\frac{\beta_{j}}{\sqrt{\frac{1-R_{Y, 2, k, k}^{2}}{N-k-1} \sqrt{\frac{1}{1-R_{j, 1,2, k}^{2}(\text { omitting } j)}}}}
$$

Where: $\quad d f=N-p-1$
Both tests will yield the same value of $t$.

The Wherry (1931) adjustment is based on defining the mean square for the residual as the sum of squares divided by the degrees of freedom rather than $\mathrm{N}-1$. This adjustment is used in SPSS and is considered an unbiased estimate of the population value.

$$
R_{\text {adjusted }}^{2}=1-\left(1-R^{2}\right) \frac{N-1}{N-p-1}
$$

Given $\mathrm{R}^{2}=.50$,
$\mathrm{N}=50, \mathrm{p}=10$
$\mathbf{R}_{\text {adjusted }}=.37$
The Stein (1960) (correlation model) adjustment is an estimate of the expected value of a series of cross validation samples where the predictors are considered random (i.e., the X's can take any value).

$$
R_{\text {adjusted }}^{2}=1-\left[\frac{N-1}{N-p-1}\right]\left[\frac{N-2}{N-p-2}\right]\left[\frac{N+1}{N}\right]\left[1-R^{2}\right]
$$

Given $\mathrm{R}^{2}=.50$,

| Data Used in the Next Example |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| A | B | c | Y |  |
| 24 | 29 | 21 | 49 |  |
| 24 | 33 | 23 | 54 |  |
| 27 | 30 | 23 | 56 |  |
| 26 | 26 | 19 | 45 |  |
| 23 | 32 | 20 | 48 |  |
| 25 | 27 | 19 | 39 |  |
| 26 | 30 | 20 | 50 |  |
| 30 | 33 | 24 | 55 |  |
| 29 | 32 | 20 | 55 |  |
| 25 | 30 | 20 | 50 |  |
| 21 | 27 | 18 | 39 |  |
| 24 | 29 | 21 | 52 |  |
| 28 | 30 | 21 | 56 |  |
| 23 | 30 | 22 | 52 |  |
| 23 | 30 | 20 | 50 |  |
| 26 | 30 | 20 | 51 |  |
| 23 | 29 | 19 | 43 |  |
| 24 | 31 | 18 | 48 |  |
| 23 | 31 | 21 | 51 |  |
| 27 | 31 | 21 | 57 | 18 |


| Correlations |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $y$ | a | b | c |  |
| Pearson | Correlation | 1.000 | 601 | . 706 | . 738 |  |
|  |  | . 601 | 1.000 | . 331 | 434 |  |
|  |  | . 706 | .331 | 1.000 | . 580 |  |
|  |  | . 738 | . 434 | . 580 | 1.000 |  |
| Sig. (1-t |  |  | . 003 | . 000 | . 000 |  |
|  |  | . 003 | . | . 077 | . 028 |  |
|  |  | . 000 | . 077 |  | . 004 |  |
|  |  | . 000 | 028 | . 004 |  |  |
| N |  | 20 | 20 | 20 | 20 |  |
|  |  | 20 | 20 | 20 | 20 |  |
|  |  | 20 | 20 | 20 | 20 |  |
|  |  | 20 | 20 | 20 | 20 |  |
| The multip | correlati | obtained | with all three | predi | s is: |  |
|  |  | Model Sum | nmary |  |  |  |
| Model | R | R Square | Adjusted R Square |  | ror of imate |  |
| 1 | $0.85864^{3}$ | 0.73727 | 0.68801 |  | 92773 |  |
| a. P | dictors: (Co | stant), c. a, b |  |  |  |  |

Often, when conducting these analyses, researchers will report the multiple correlation, and the regression coefficients.

a. Dependent Variable: y

In the present example, all three regression coefficients are significant. This does not mean they are significant predictors (this information is contained in the correlation matrix). It means only that given the criterion and these three variables each one adds significantly to prediction. Add or subtract one or more predictors or change the dependent variable and the results can change drastically. This is because the multiple correlation is made up of more than the information given in the regression coefficients ${ }_{2}$

## On the interpretation of regression coefficients

Unstandardized regression coefficients are the weights applied to the original measures. As such they are expressed in the unit of measurement of the variable, and thus are not directly comparable. They indicate the amount of change in the dependent variable for a unit change in the predictor. Tests of significance of regression coefficients are performed on these coefficients.

Standardized regression coefficients are the weights applied to the standardized measures and define the amount of change in the standardized dependent variable for a unit change in the standardized predictor. They are unitless, and are weights of variables that have a mean of 0 , and a standard deviation of 1 . They are not, however, directly comparable. Given the relationships among the variables, it is possible for one Beta to be larger than another, though the second may be significant ${ }_{92}$ and the first not.

## Venn Diagrams and the Dimensionality of Multiple Correlation

Consider a three predictor problem, A, B, and C. The following Venn diagram shows the sources of variance overlapping the criterion. (What about the other stuff?)


This can be demonstrated by computing the multiple correlations with each pair of predictors and calculating the component accounted for by each segment in Slide 23.

Model Summary

a. Predictors: (Constant). c, b

Model Summary

| Model | $R$ | R Square | Adjusted $R$ <br> Square | Std. Error of <br> the Estimate |
| :---: | :---: | :---: | :---: | :---: |
| 1 | $0.80110^{\text {a }}$ | 0.64176 | 0.59962 | 3.31663 |

a. Predictors: (Constant), a, c

| Model | $R$ | R Square | Adjusted $R$ <br> Square | Std. Error of <br> the Estimate |
| :---: | :---: | :---: | :---: | :---: |
| 1 | $0.80618^{\mathrm{a}}$ | 0.64993 | 0.60875 | 3.27859 |
|  |  |  |  |  |
| a. Predictors: (Constant), b, a |  |  |  |  |

a. Predictors: (Constant), b, a


Contributions to the Squared Multiple Correlation
$\mathrm{AB}=.02069$
$\mathrm{Note}:$
$\mathrm{r}_{\mathrm{AY}}=.07669+.07513+.18888+.02069=.36139=.60116^{2}=\mathrm{r}^{2}{ }_{\mathrm{AY}}$
$\mathrm{r}^{2}{ }_{\mathrm{BY}}=.09551+.02069+.18888+.19303=.49811=.70577^{2}=\mathrm{r}^{2}{ }_{\mathrm{BY}}$
$\mathrm{r}^{2}{ }_{\mathrm{CY}}=.08734+.07513+.18888+.19303=.54438=73782^{2}=\mathrm{r}^{2}{ }_{\mathrm{CY}}$

## Comments on these values

Segments 1,3 , and 7 are squared semipartial multiple correlations, thus they are always positive. Moreover, their tests of significance are identical to tests of significance of the corresponding regression coefficients.
Segments $2,4,5$, and 6 are residuals and thus can be positive or negative. There are no obvious tests of significance, but the segments sum to $R^{2}$, thus it is possible to estimate the proportion that each contributes to $\mathrm{R}^{2}$.

| Segment 1 | .07669/.73727, ..10.40\% | Unique A |  |
| :---: | :---: | :---: | :---: |
| Segment 3 | .09551/.73727 ... 12.95\% | Unique B | 35.2\% |
| Segment 7 | .08734/.73727, ..11.85\% | Unique C |  |
| Segment 4 | .02069/.73727 ... 2.81\% | Unique AB |  |
| Segment 6 | .07513/.73727, ...10.19\% | Unique AC |  |
| Segment 2 | .19303/.73727, ...26.18\% | Unique BC |  |
| Segment 5 | .18888/.73727, ...25.62\% | Unique ABC |  |
|  | Total 100.00\% |  |  | 27

## Methods of Indirect Entry

Methods of indirect entry use a series of steps followed by the computer to determine the least number of predictors that give almost as good a level of prediction as the full number of predictors. These methods are good for the purpose intended, but are not of any value for interpretation.

## Stepwise Inclusion

Step 1. The predictor that correlates highest with the criterion is the first predictor.
Step 2. Partial correlations, removing the effects of the first predictor, are contrasted, and the predictor with the highest partial correlation is the next predictor.

Step 3. The regression equation is determined, and each regression coefficient is tested for significance. If any are not significant, they are removed.

Step 4. Calculate the partial correlations removing effects of all predictors in the regression equation after Step 3. If none are significant, stop; otherwise go to step 3.

## Forward Inclusion

The computer follows the previous steps except that the regression
coefficients are not tested for significance after each equation is determined.
As a result, some coefficients may not be significant in the final equation

Backward Elimination

Step 1. All variables are entered into the equation, and R is calculated.

Step 2. Test regression coefficients for significance. If any are not significant remove the predictor that has the highest alpha.

Step 3. Calculate new $R$ and test regression coefficients for significance. If any are not significant, remove the one with the highest alpha. Repeat until all regression coefficients are significant.

