

## PSY 9556B (Feb 5) Latent Growth Modeling

- “Fixed” and “random” word confusion
- Simplest LGM – knowing how to calculate dfs
- How many time points needed?
- Power, sample size
- Nonlinear growth – quadratic
- Nonlinear growth freeing loadings
- Piecewise models
- Linear growth (different ways of scaling time)
- Associative LGM
- Higher order LGM: curve-of-factors model
- Conditional models (time-invariant, time-variant)
- Multiple groups (group covariate or multiple-groups analysis)
- Similarity between LGM and MLM
- When to use LGM, when to use MLM

## “Fixed” and “Random” word clarification

- **Fixed and random effects** in MLM and LGM
  - Fixed effect: single values that estimate of population values
    - (e.g., a regression coefficient, a mean intercept or slope in LGM, MLM)
  - Random effect: provide information about the variation in the regression coefficient or intercept parameters across the clustering units
    - (e.g., variance of intercepts and slopes in LGM or MLM)
- **Fixed and random factors** in ANOVA
  - Fixed factor levels chosen apriori
  - Random factor: no particular interest in the levels; chosen at random
    - Best example of a random factor: persons
    - Repeated measures design
    - At least two observations nested within persons
    - Persons as a random factor is also referred to as the clustering unit or in MLM
- Mixed models
  - One fixed factor crossed with one random factor (e.g., split plot ANOVA)

## Number of Parameters and Degrees of Freedom

### Example: 2 time points (linear)

#### Parameters and dfs

Elements:

$$(v(v+3))/2 = (2*5)/2 = 5$$

Parameters:

**2 residuals (2 time points):** left-over variance not explained by latent variables

**1 mean intercept:** the mean start-point of individual trajectories

**1 mean slope:** the mean slope (e.g., growth/learning/decrease) of individual trajectories

**1 variance of the intercepts:** variation in individual start-points

**1 variance of the slopes:** variation in individual slopes

**1 correlation between intercept and slope**

(note, indicator intercepts fixed at 0)

Total parameters = 7

dfs = to many parameters

## Number of Parameters and Degrees of Freedom

### Example: 3 time points (linear)

#### Parameters and dfs

Elements:

$$(v(v+3))/2 = (3*6)/2 = 9$$

Parameters:

**3 residuals (3 time points):** left-over variance not explained by latent variables

**1 mean intercept:** the mean start-point of individual trajectories

**1 mean slope:** the mean slope (e.g., growth/learning/decrease) of individual trajectories

**1 variance of the intercepts:** variation in individual start-points

**1 variance of the slopes:** variation in individual slopes

**1 correlation between intercept and slope**

(note, indicator intercepts fixed at 0)

Total parameters = 8

$$\text{dfs} = 9 - 8 = 1$$

## Number of Parameters and Degrees of Freedom

### Example: 4 time points (linear + quadratic)

#### Parameters and dfs

Elements:

$$(4(4+3))/2 = (4*7)/2 = 14$$

Parameters:

**4 residuals (4 time points):** left-over variance not explained by latent variables

**1 mean intercept:** the mean start-point of individual trajectories

**1 mean slope:** the mean slope (e.g., growth/learning/decrease) of individual trajectories

**1 mean quadratic component**

**1 variance of the intercepts:** variation in individual start-points

**1 variance of the slopes:** variation in individual slopes

**1 variance of quadratic component**

**3 correlations between intercept, slope, quadratic**

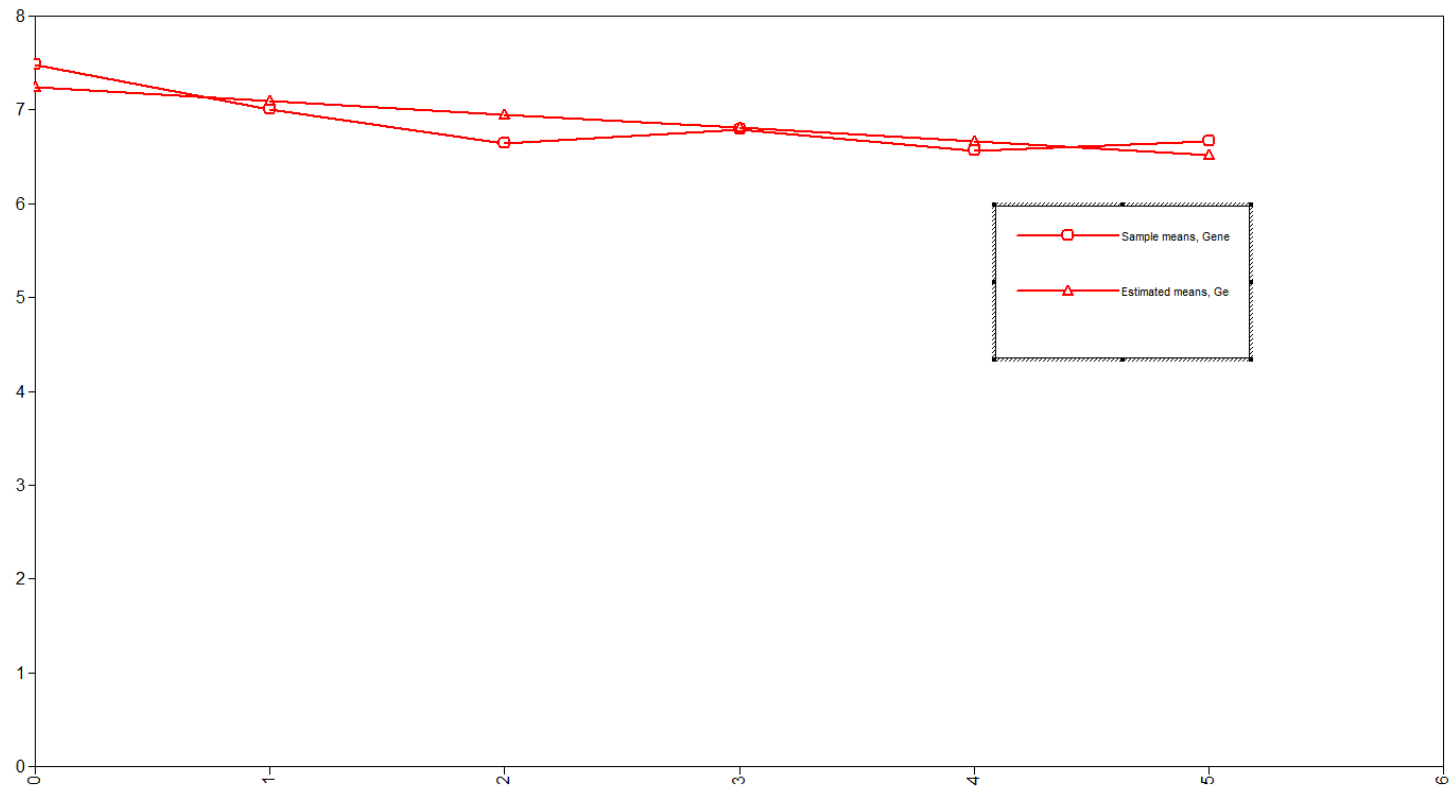
(note, indicator intercepts fixed at 0)

Total parameters = 13

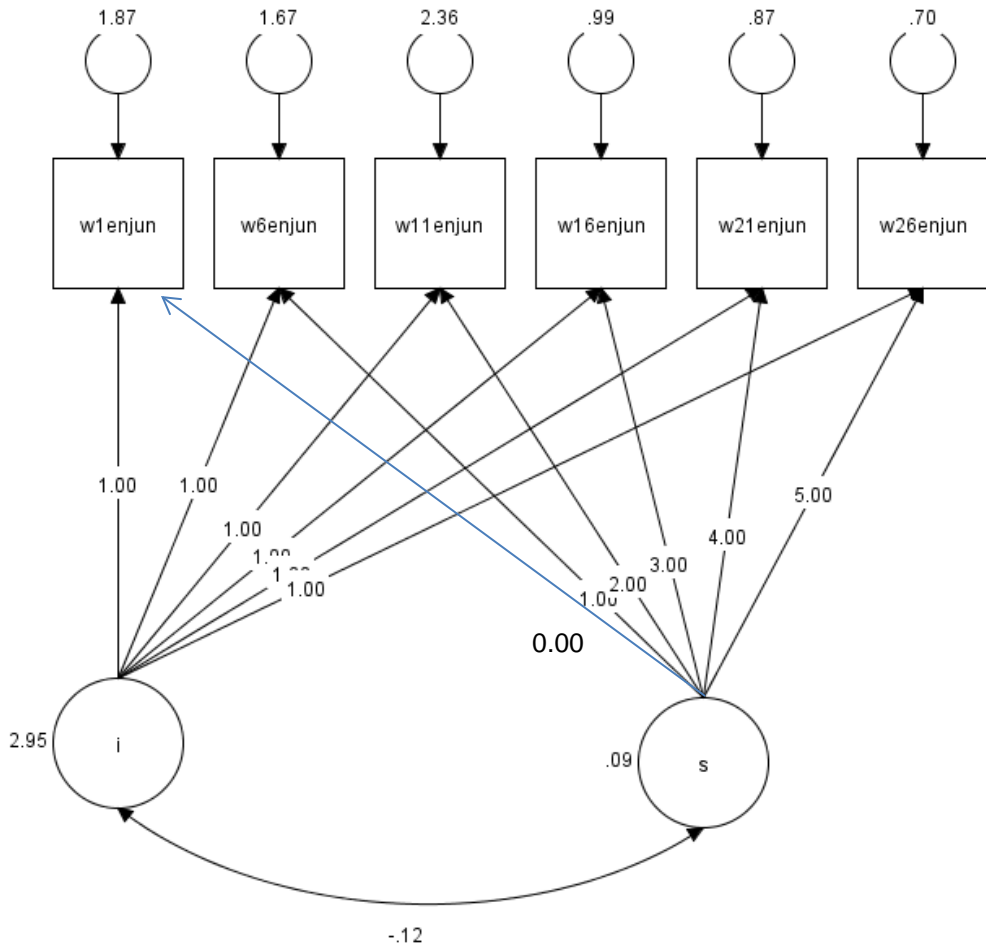
$$\text{dfs} = 14 - 13 = 1$$

## Example of a LGM with Five Time Points

```
USEVARIABLES ARE w1enjun w6enjun w11enjun w16enjun w21enjun w26enjun;  
model:  
I S | w1enjun@0 w6enjun@1 w11enjun@2 w16enjun@3 w21enjun@4 w26enjun@5;  
plot:  
type=plot3;  
series = w1enjun(0) w6enjun(1) w11enjun(2) w16enjun(3) w21enjun(4) w26enjun(5);  
output: sampstat residual stdyx tech4 modindices;
```



## Example of a LGM with Five Time Points



In this figure you should also indicate the means of the latent variables: Intercept = 2.948 and Slope = 0.094

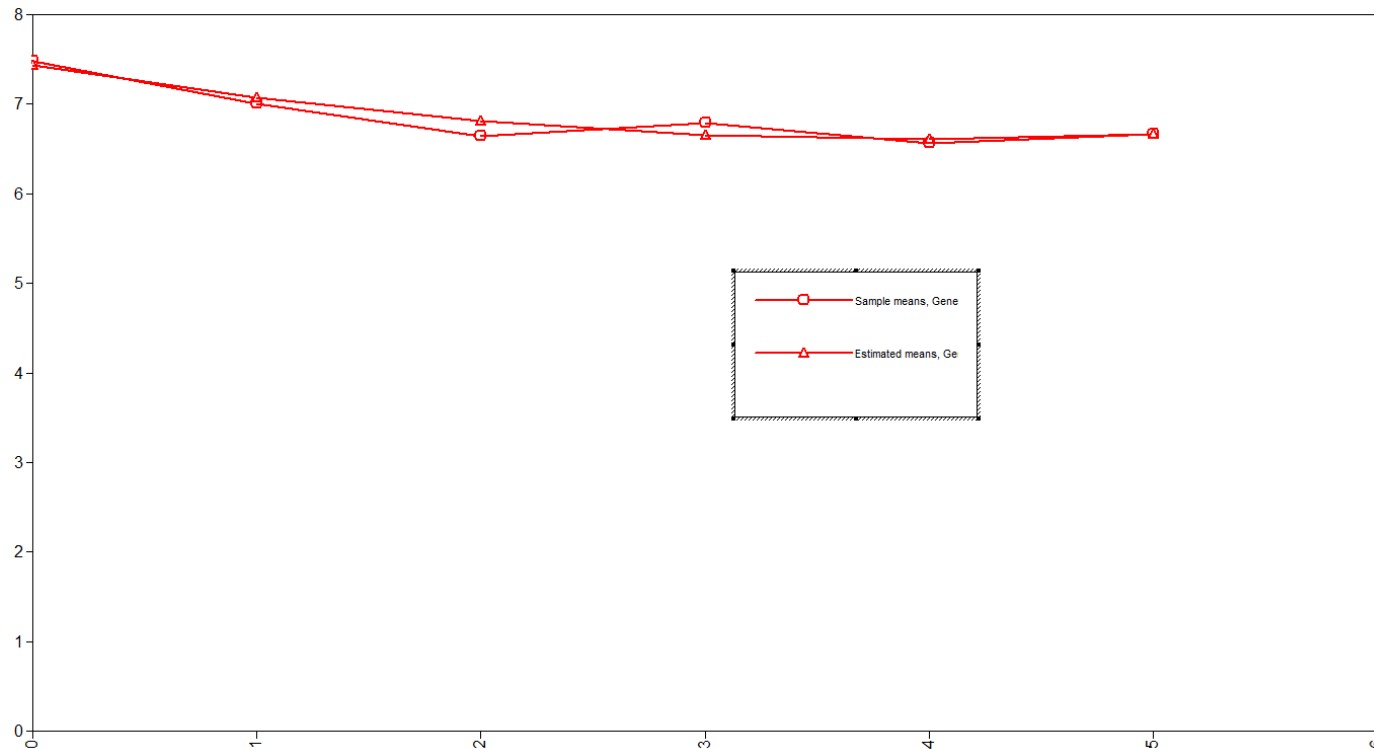
## Example of a LGM with Five Time Points

			Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Number of Free Parameters	11					
Loglikelihood						
H0 Value	-3779.045					
H1 Value	-3745.014					
Information Criteria						
Akaike (AIC)	7580.091					
Bayesian (BIC)	7624.349					
Sample-Size Adjusted BIC	7589.443					
(n* = (n + 2) / 24)						
Chi-Square Test of Model Fit						
Value	68.063					
Degrees of Freedom	16					
P-Value	0.0000					
RMSEA (Root Mean Square Error Of Approximation)						
Estimate	0.089					
90 Percent C.I.	0.068	0.111				
Probability RMSEA <= .05	0.002					
CFI/TLI						
CFI	0.957					
TLI	0.960					
Chi-Square Test of Model Fit for the Baseline Model						
Value	1239.482					
Degrees of Freedom	15					
P-Value	0.0000					
SRMR (Standardized Root Mean Square Residual)						
Value	0.078					

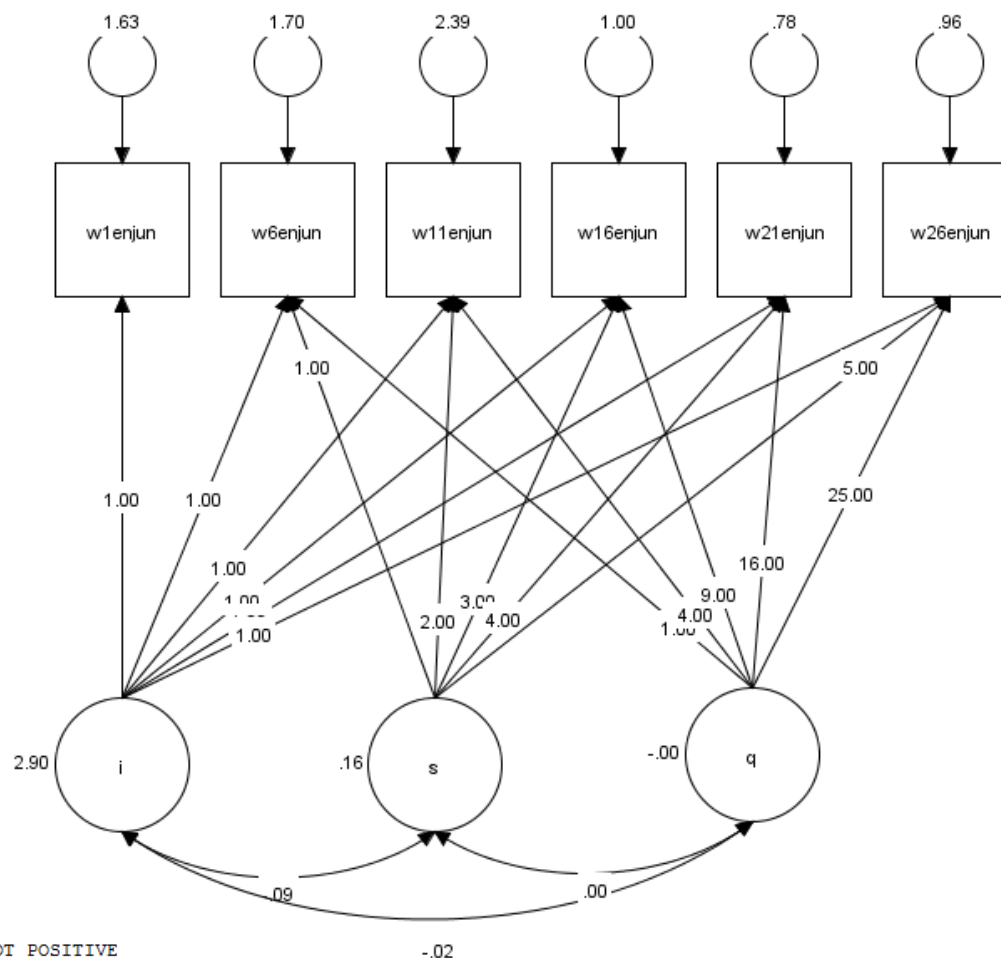


## Example of a LGM with Five Time Points Adding a Quadratic Component

```
USEVARIABLES ARE wlenjun w6enjun w11enjun w16enjun w21enjun w26enjun;  
model:  
I S Q| wlenjun@0 w6enjun@1 w11enjun@2 w16enjun@3 w21enjun@4 w26enjun@5;  
plot:  
type=plot3;  
series = wlenjun(0) w6enjun(1) w11enjun(2) w16enjun(3) w21enjun(4) w26enjun(5);  
output: sampstat residual stdyx tech4 modindices;
```



## Example of a LGM with Five Time Points Adding a Quadratic Component

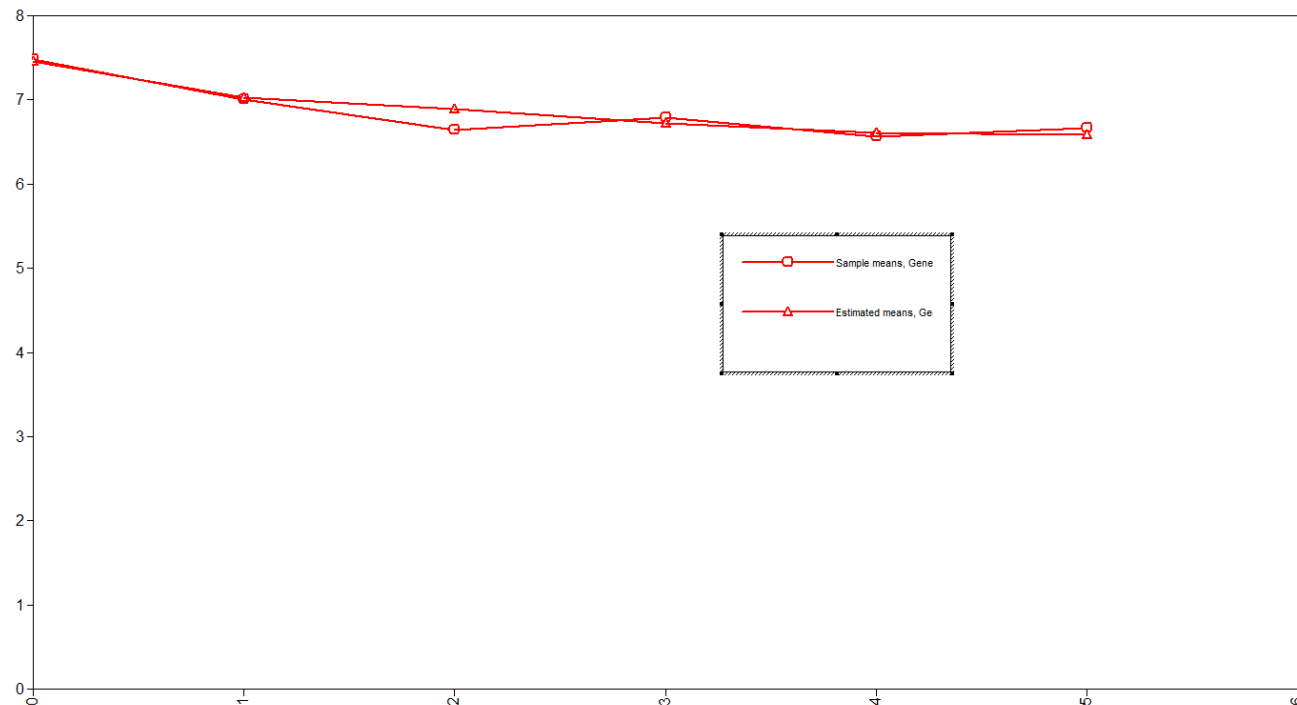


THE MODEL ESTIMATION TERMINATED NORMALLY

WARNING: THE LATENT VARIABLE COVARIANCE MATRIX (PSI) IS NOT POSITIVE DEFINITE. THIS COULD INDICATE A NEGATIVE VARIANCE/RESIDUAL VARIANCE FOR A LATENT VARIABLE, A CORRELATION GREATER OR EQUAL TO ONE BETWEEN TWO LATENT VARIABLES, OR A LINEAR DEPENDENCY AMONG MORE THAN TWO LATENT VARIABLES. CHECK THE TECH4 OUTPUT FOR MORE INFORMATION. PROBLEM INVOLVING VARIABLE Q.

## Example of a LGM with Five Time Points Time Points Free

```
USEVARIABLES ARE w1enjun w6enjun w11enjun w16enjun w21enjun w26enjun;  
model:  
I S | w1enjun@0 w6enjun@1 w11enjun* w16enjun* w21enjun* w26enjun*5;  
plot:  
type=plot3;  
series = w1enjun(0) w6enjun(1) w11enjun(2) w16enjun(3) w21enjun(4) w26enjun(5);  
output: sampstat residual stdyx tech4 modindices;
```

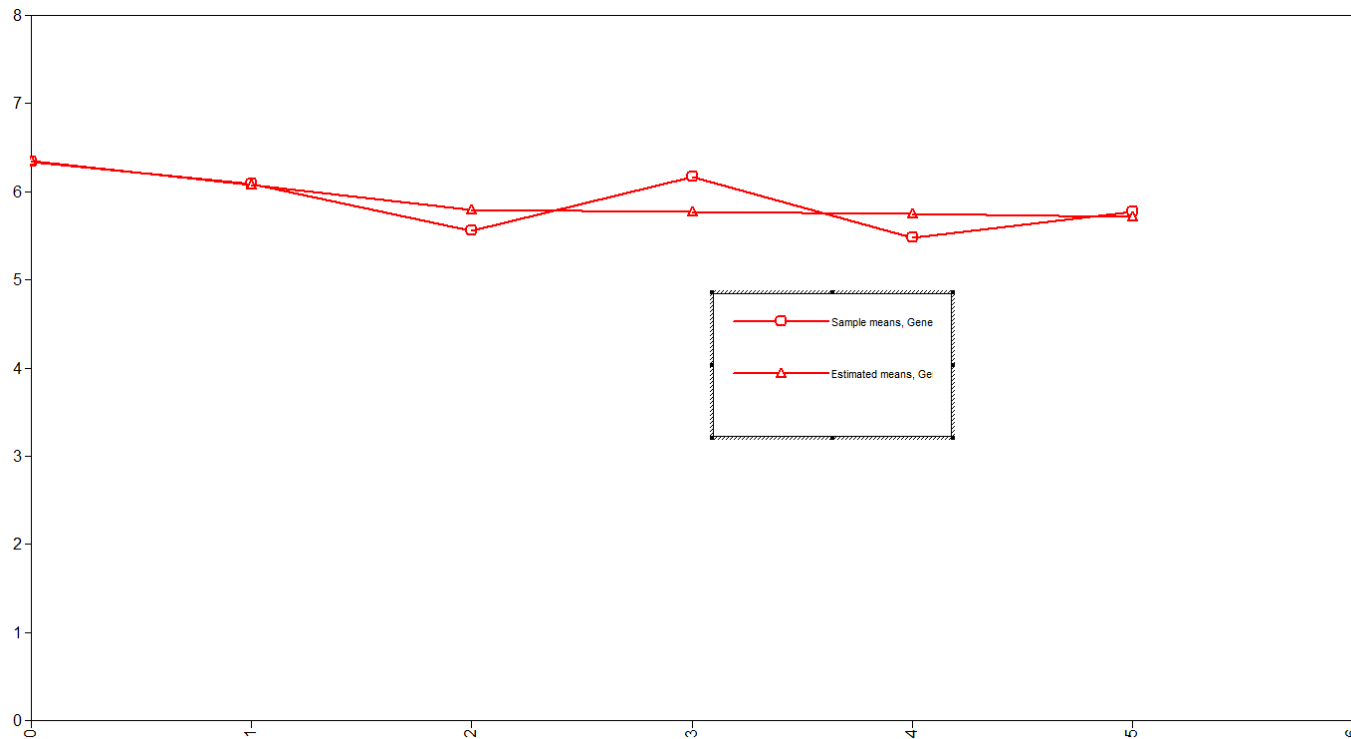


## Example of a LGM with Five Time Points Time Points Free

			Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Number of Free Parameters		15				
Loglikelihood			I	I		
				W1ENJUN	1.000	0.000
				W6ENJUN	1.000	0.000
	H0 Value	-3764.199		W11ENJUN	1.000	0.000
	H1 Value	-3745.014		W16ENJUN	1.000	0.000
				W21ENJUN	1.000	0.000
				W26ENJUN	1.000	0.000
Information Criteria			S	I		
	Akaike (AIC)	7558.398		W1ENJUN	0.000	0.000
	Bayesian (BIC)	7618.749		W6ENJUN	1.000	0.000
	Sample-Size Adjusted BIC	7571.151		W11ENJUN	1.304	0.173
	(n* = (n + 2) / 24)			W16ENJUN	1.706	0.233
				W21ENJUN	1.954	0.287
				W26ENJUN	2.013	0.289
Chi-Square Test of Model Fit			S	WITH		
	Value	38.370	I		-0.687	0.288
	Degrees of Freedom	12				
	P-Value	0.0001	Means			
			I		7.452	0.108
			S		-0.429	0.087
RMSEA (Root Mean Square Error Of Approximation)						
	Estimate	0.073	Intercepts			
	90 Percent C.I.	0.048	0.099	W1ENJUN	0.000	0.000
	Probability RMSEA <= .05	0.065		W6ENJUN	0.000	0.000
				W11ENJUN	0.000	0.000
				W16ENJUN	0.000	0.000
				W21ENJUN	0.000	0.000
				W26ENJUN	0.000	0.000
CFI/TLI						
	CFI	0.978	Variances			
	TLI	0.973	I		3.523	0.444
			S		0.723	0.291
Chi-Square Test of Model Fit for the Baseline Model						
	Value	1239.482	Residual Variances			
	Degrees of Freedom	15		W1ENJUN	0.935	0.374
	P-Value	0.0000		W6ENJUN	1.869	0.177
				W11ENJUN	2.435	0.218
				W16ENJUN	0.939	0.101
				W21ENJUN	0.774	0.111
				W26ENJUN	0.910	0.124
SRMR (Standardized Root Mean Square Residual)						
	Value	0.030				

## Example of a LGM with Five Time Points Piecewise

```
USEVARIABLES ARE w1enjcr w6enjcr w11enjcr w16enjcr w21enjcr w26enjcr;  
model:  
I S1 | w1enjcr@0 w6enjcr@1 w11enjcr@2 w16enjcr@2 w21enjcr@2 w26enjcr@2;  
I S2 | w1enjcr@0 w6enjcr@0 w11enjcr@0 w16enjcr@1 w21enjcr@2 w26enjcr@3;  
plot:  
type=plot3;  
series = w1enjcr(0) w6enjcr(1) w11enjcr(2) w16enjcr(3) w21enjcr(4) w26enjcr(5);  
output: sampstat residual stdyx tech4 modindices;
```



## Example of a LGM with Five Time Points Piecewise

### MODEL FIT INFORMATION

Number of Free Parameters 15

### Loglikelihood

H0 Value -4038.467  
H1 Value -3997.736

### Information Criteria

Akaike (AIC) 8106.934  
Bayesian (BIC) 8167.249  
Sample-Size Adjusted BIC 8119.651  
( $n^* = (n + 2) / 24$ )

### Chi-Square Test of Model Fit

Value 81.462  
Degrees of Freedom 12  
P-Value 0.0000

### RMSEA (Root Mean Square Error Of Approximation)

Estimate 0.119  
90 Percent C.I. 0.095 0.144  
Probability RMSEA  $\leq$  .05 0.000

### CFI/TLI

CFI 0.902  
TLI 0.877

### Chi-Square Test of Model Fit for the Baseline Model

Value 722.674  
Degrees of Freedom 15  
P-Value 0.0000

### SRMR (Standardized Root Mean Square Residual)

Value 0.074

Poor model

## Power in LGM

- Fan (2003) showed that LGM consistently showed higher statistical power for detecting group differences in the linear growth slope than repeated measures ANOVA.
- Study done with five time points
- More research needed on the impact of number of time points

Fan, X. (2003). Power of latent growth modeling for detecting group differences in linear growth trajectory parameters. *Structural Equation Modeling: A Multidisciplinary Journal*, 10, 380-400.

# Alternative Ways of Scaling Slope: Drinking Example (8 time points)

## MODEL FIT INFORMATION

Number of Free Parameters 13

## Loglikelihood

H0 Value -7809.924

H1 Value -7736.326

## Information Criteria

Akaike (AIC) 15645.849

Bayesian (BIC) 15695.927

Sample-Size Adjusted BIC 15654.687

( $n^* = (n + 2) / 24$ )

## Chi-Square Test of Model Fit

Value 147.197

Degrees of Freedom 31

P-Value 0.0000

## RMSEA (Root Mean Square Error Of Approximation)

Estimate 0.104

90 Percent C.I. 0.087 0.121

Probability RMSEA  $\leq$  .05 0.000

## CFI/TLI

CFI 0.942

TLI 0.948

## Chi-Square Test of Model Fit for the Baseline Model

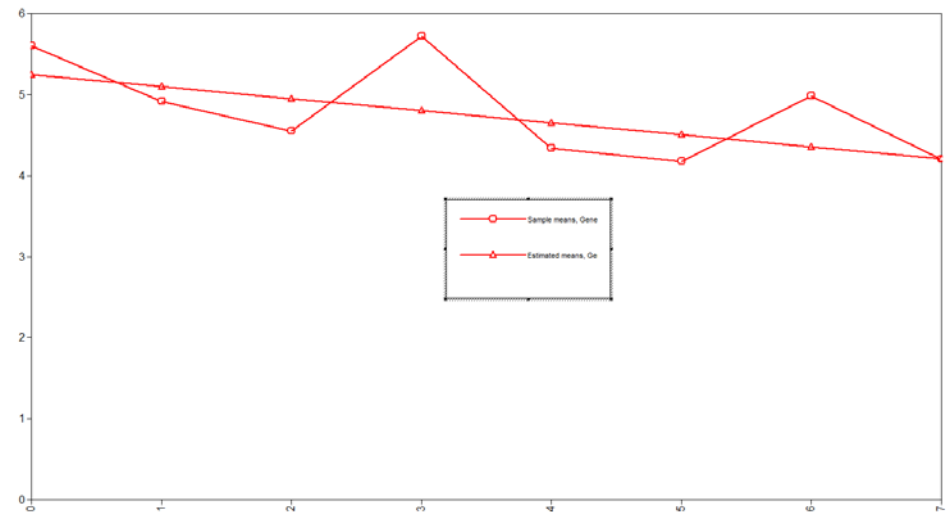
Value 2027.634

Degrees of Freedom 28

P-Value 0.0000

## SRMR (Standardized Root Mean Square Residual)

Value 0.059





## Alternative Ways of Scaling Slope: Drinking Example (8 time points)

```
usevariables are d1 d2 d3 d4 d5 d6 d7 d8;
model: i s | d1@0 d2@0.143 d3@0.286 d4@0.429 d5@0.572 d6@0.715 d7@0.858 d8@1;
_
```

S	WITH				
I		-13.709	2.624	-5.225	0.000
Means					
I		5.246	0.351	14.941	0.000
S		-1.038	0.348	-2.978	0.003
Variances					
I		36.513	3.251	11.230	0.000
S		24.601	3.386	7.266	0.000

```
usevariables are d1 d2 d3 d4 d5 d6 d7 d8;
model: i s | d1@0 d2@1 d3@2 d4@3 d5@4 d6@5 d7@6 d8@7;
```

S	WITH				
I		-1.959	0.375	-5.225	0.000
Means					
I		5.245	0.351	14.941	0.000
S		-0.148	0.050	-2.978	0.003
Variances					
I		36.511	3.251	11.230	0.000
S		0.502	0.069	7.266	0.000

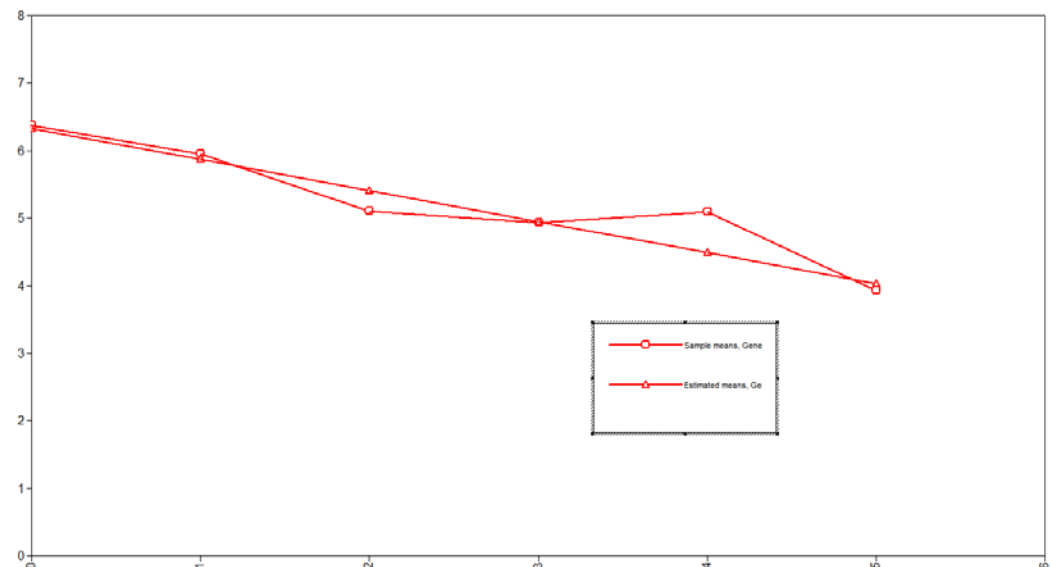
```
usevariables are d1 d2 d3 d4 d5 d6 d7 d8;
model: i s | d1@-7 d2@-6 d3@-5 d4@-4 d5@-3 d6@-2 d7@-1 d8@0;
```

S	WITH				
I		1.557	0.359	4.332	0.000
Means					
I		4.208	0.340	12.385	0.000
S		-0.148	0.050	-2.978	0.003
Variances					
I		33.698	3.047	11.060	0.000
S		0.502	0.069	7.266	0.000

## Associative LGM (i.e., Trajectory of Two Different Variables)

I will investigate whether the slope in enjoying courses is associated with the slope in drinking. Here is a LGM of drinking only:

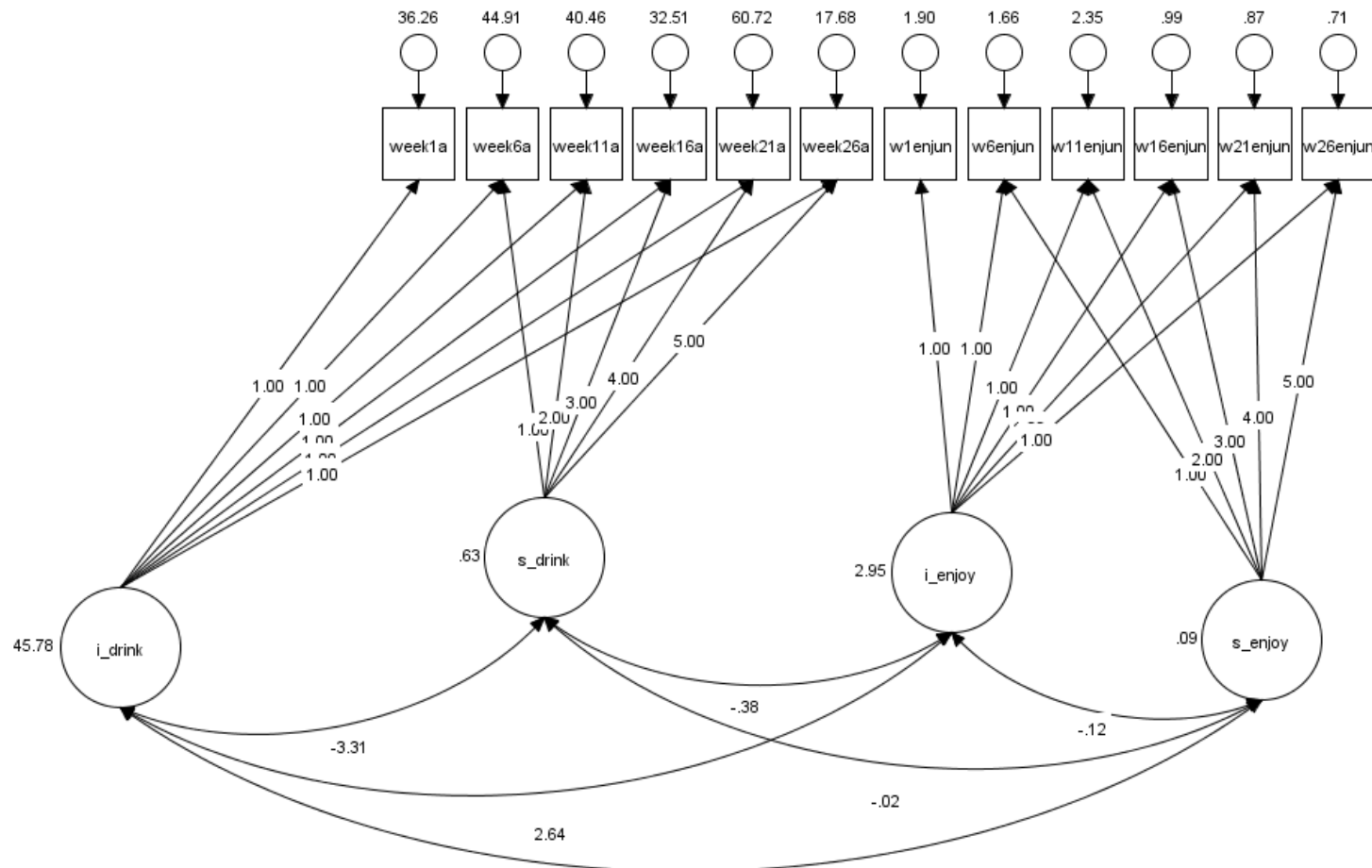
Number of Free Parameters	11
Loglikelihood	
H0 Value	-6995.646
H1 Value	-6952.667
Information Criteria	
Akaike (AIC)	14013.292
Bayesian (BIC)	14057.550
Sample-Size Adjusted BIC	14022.645
( $n^* = (n + 2) / 24$ )	
Chi-Square Test of Model Fit	
Value	85.958
Degrees of Freedom	16
P-Value	0.0000
RMSEA (Root Mean Square Error Of Approximation)	
Estimate	0.103
90 Percent C.I.	0.082 0.125
Probability RMSEA <= .05	0.000
CFI/TLI	
CFI	0.888
TLI	0.895
Chi-Square Test of Model Fit for the Baseline Model	
Value	642.245
Degrees of Freedom	15
P-Value	0.0000
SRMR (Standardized Root Mean Square Residual)	
Value	0.066



S	WITH				
I		-3.351	0.969	-3.459	0.001
Means					
I		6.328	0.406	15.568	0.000
S		-0.459	0.083	-5.551	0.000
Variances					
I		45.801	5.084	9.009	0.000
S		0.650	0.288	2.260	0.024

## Associative LGM (i.e., Trajectory of Two Different Variables)

```
USEVARIABLES ARE week1a week6a week11a week16a week21a week26a
w1enjun w6enjun w11enjun w16enjun w21enjun w26enjun;
model:
I_drink S_drink | week1a@0 week6a@1 week11a@2 week16a@3 week21a@4 week26a@5;
I_enjoy S_enjoy | w1enjun@0 w6enjun@1 w11enjun@2 w16enjun@3 w21enjun@4 w26enjun@5;
output: sampstat residual stdyx tech4 modindices;
```



## Associative LGM (i.e., Trajectory of Two Different Variables)

### Loglikelihood

H0 Value -10767.889  
H1 Value -10665.836

### Information Criteria

Akaike (AIC) 21587.778  
Bayesian (BIC) 21692.388  
Sample-Size Adjusted BIC 21609.884  
(n\* = (n + 2) / 24)

### Chi-Square Test of Model Fit

Value 204.105  
Degrees of Freedom 64  
P-Value 0.0000

### RMSEA (Root Mean Square Error Of Approximation)

Estimate 0.073  
90 Percent C.I. 0.062 0.084  
Probability RMSEA <= .05 0.000

### CFI/TLI

CFI 0.925  
TLI 0.923

### Chi-Square Test of Model Fit for the Baseline Model

Value 1945.417  
Degrees of Freedom 66  
P-Value 0.0000

### SRMR (Standardized Root Mean Square Residual)

Value 0.066

### Means

I_DRINK	6.309	0.407	15.514	0.000
S_DRINK	-0.454	0.083	-5.485	0.000
I_ENJOY	7.240	0.100	72.342	0.000
S_ENJOY	-0.144	0.022	-6.408	0.000

### Variances

I_DRINK	45.779	5.085	9.003	0.000
S_DRINK	0.631	0.287	2.201	0.028
I_ENJOY	2.949	0.290	10.181	0.000
S_ENJOY	0.093	0.014	6.705	0.000

### S\_DRINK WITH

I_DRINK	-0.616	0.090	-6.833	0.000
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### I\_ENJOY WITH

I_DRINK	0.228	0.069	3.307	0.001
S_DRINK	-0.277	0.133	-2.078	0.038

### S\_ENJOY WITH

I_DRINK	-0.013	0.089	-0.145	0.884
S_DRINK	-0.087	0.146	-0.599	0.549
I_ENJOY	-0.220	0.081	-2.701	0.007

## One Trajectory with Time Varying Covariates

I will investigate the previous trajectory of drinking and add time varying covariates of Enjoyed University. Recall that in the previous example, Enjoyed University was modelled as a separate trajectory. In the present example, I ignore the growth factor of Enjoyed University.

```
model:
I_drink S_drink | week1a@0 week6a@1 week11a@2 week16a@3 week21a@4 week26a@5;
week1a on w1enjun;
week6a on w6enjun;
week11a on w11enjun;
week16a on w16enjun;
week21a on w21enjun;
week26a on w26enjun;
I_drink S_drink with w1enjun w6enjun w11enjun w16enjun w21enjun w26enjun;
output: sampstat residual stdyx tech4 modindices;
```

# One Trajectory with Time Varying Covariates

Number of Free Parameters 56

Loglikelihood

H0 Value -10728.518  
H1 Value -10665.836

Not a good model;  
covariate do not add much

Information Criteria

Akaike (AIC) 21569.035  
Bayesian (BIC) 21794.348  
Sample-Size Adjusted BIC 21616.648  
( $n^* = (n + 2) / 24$ )

Chi-Square Test of Model Fit

Value 125.363  
Degrees of Freedom 34  
P-Value 0.0000

RMSEA (Root Mean Square Error Of Approximation)

Estimate 0.081  
90 Percent C.I. 0.066 0.096  
Probability RMSEA <= .05 0.000

CFI/TLI

CFI 0.861  
TLI 0.791

Chi-Square Test of Model Fit for the Baseline Model

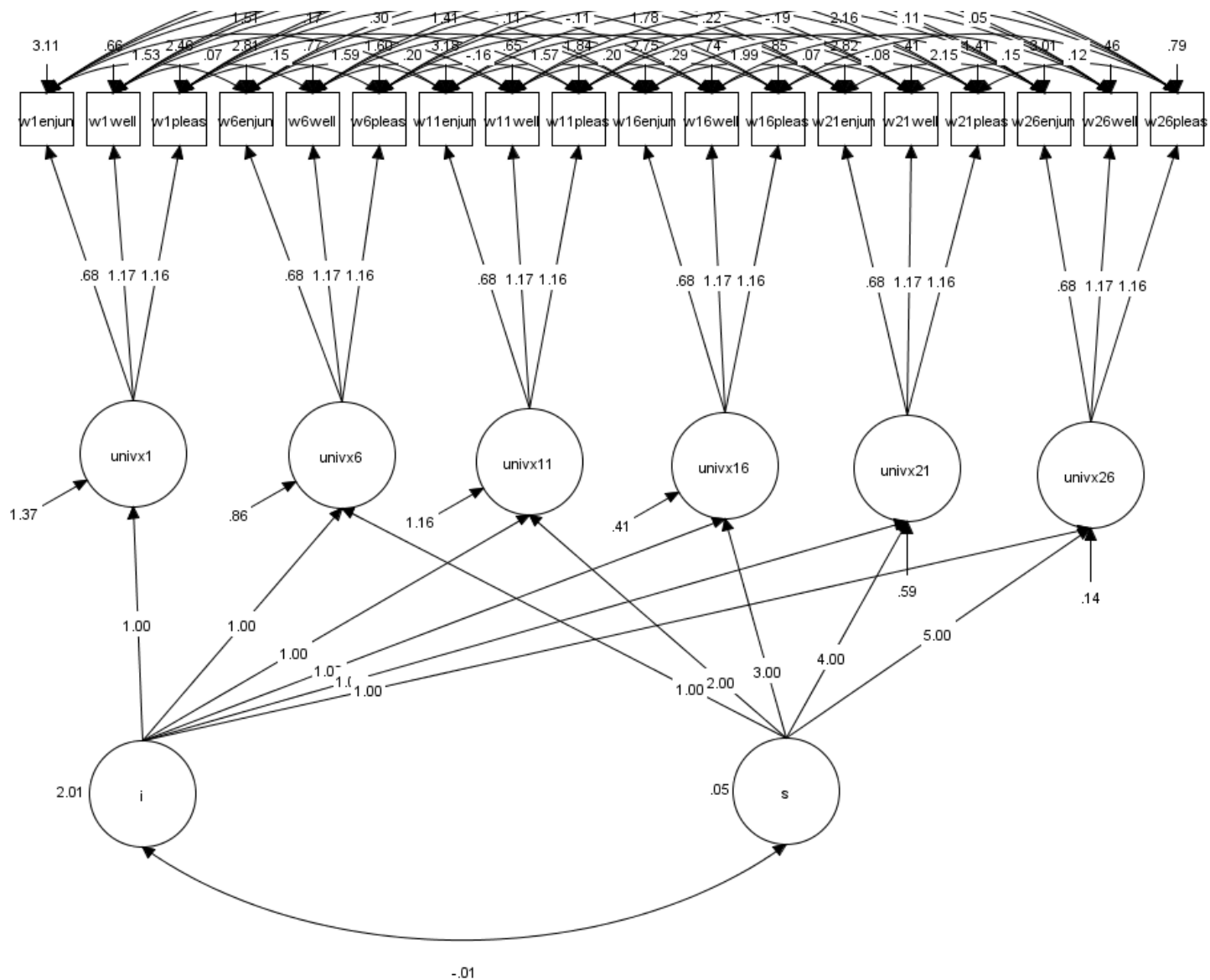
Value 705.935  
Degrees of Freedom 51  
P-Value 0.0000

SRMR (Standardized Root Mean Square Residual)

Value 0.045

WEEK1A ON				
W1ENJUN	0.151	0.225	0.671	0.502
WEEK6A ON				
W6ENJUN	0.228	0.179	1.274	0.203
WEEK11A ON				
W11ENJUN	0.145	0.159	0.914	0.361
WEEK16A ON				
W16ENJUN	0.265	0.182	1.457	0.145
WEEK21A ON				
W21ENJUN	0.336	0.255	1.320	0.187
WEEK26A ON				
W26ENJUN	0.285	0.327	0.874	0.382
I_DRINK WITH				
W1ENJUN	1.101	1.223	0.900	0.368
W6ENJUN	2.469	1.193	2.069	0.039
W11ENJUN	2.848	1.265	2.251	0.024
W16ENJUN	2.331	1.122	2.079	0.038
W21ENJUN	2.411	1.083	2.227	0.026
W26ENJUN	1.830	1.154	1.587	0.113
S_DRINK WITH				
W1ENJUN	-0.282	0.324	-0.872	0.383
W6ENJUN	-0.540	0.337	-1.602	0.109
W11ENJUN	-0.541	0.333	-1.624	0.104
W16ENJUN	-0.599	0.364	-1.648	0.099
W21ENJUN	-0.654	0.358	-1.829	0.067
W26ENJUN	-0.666	0.412	-1.619	0.106
I_DRINK	-3.332	1.000	-3.332	0.001

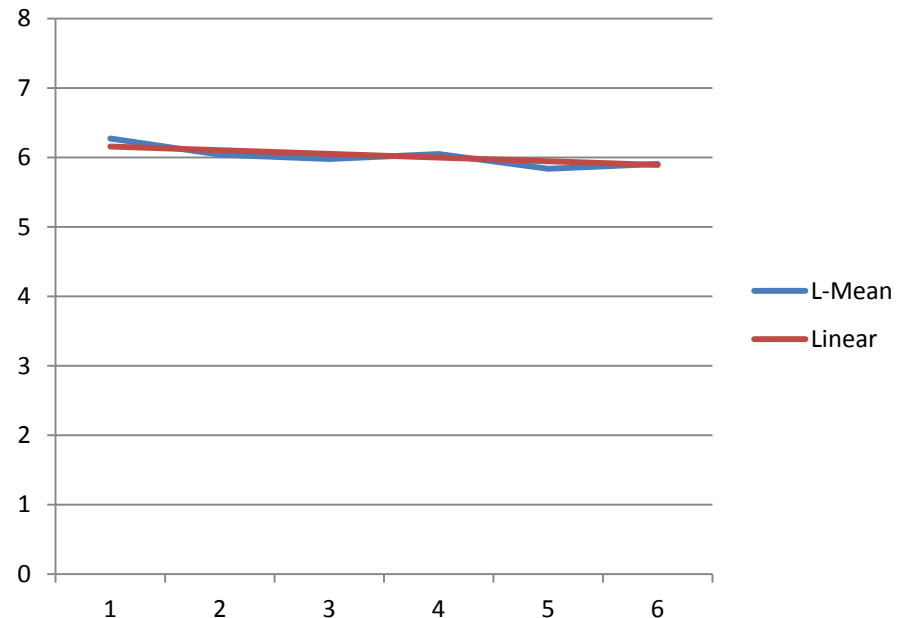
## Higher Order LGM: Curve of Factor Model



## Higher Order LGM: Curve of Factor Model

Steps:

- Allow residuals of indicators to correlate across time
- Test for strong invariance (loadings and intercepts across time points)
- Constrain loadings and intercepts to equality across time points
- Fix intercepts of the level-1 factors (i.e.,  $\text{univx1}$   $\text{univx2}$ ...in this example) at 0
- In Mplus, the intercept growth factor is fixed at 0; free this parameter





## Higher Order LGM: Curve of Factor Model

### MODEL:

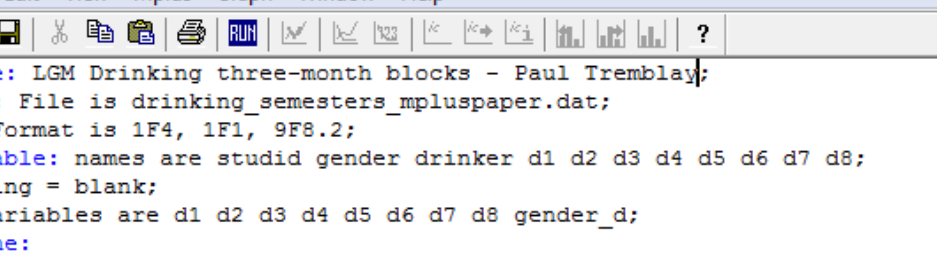
```
univx1 by w1enjun* (L1)
w1well (L2)
w1pleas (L3);
univx6 by w6enjun* (L1)
w6well (L2)
w6pleas (L3);
univx11 by w11enjun* (L1)
w11well (L2)
w11pleas (L3);
univx16 by w16enjun* (L1)
w16well (L2)
w16pleas (L3);
univx21 by w21enjun* (L1)
w21well (L2)
w21pleas (L3);
univx26 by w26enjun* (L1)
w26well (L2)
w26pleas (L3);
[w1enjun] (T1);
[w1well] (T2);
[w1pleas] (T3);
[w6enjun] (T1);
[w6well] (T2);
[w6pleas] (T3);
[w11enjun] (T1);
[w11well] (T2);
[w11pleas] (T3);
[w16enjun] (T1);
[w16well] (T2);
[w16pleas] (T3);
[w21enjun] (T1);
[w21well] (T2);
[w21pleas] (T3);
[w26enjun] (T1);
[w26well] (T2);
[w26pleas] (T3);

w1pleas with w6pleas w11pleas w16pleas w21pleas w26pleas;
w6pleas with w11pleas w16pleas w21pleas w26pleas;
w11pleas with w16pleas w21pleas w26pleas;
w16pleas with w21pleas w26pleas;
w21pleas with w26pleas;
w1well with w6well w11well w16well w21well w26well;
w6well with w11well w16well w21well w26well;
w11well with w16well w21well w26well;
w16well with w21well w26well;
w21well with w26well;
w1enjun with w6enjun w11enjun w16enjun w21enjun w26enjun;
w6enjun with w11enjun w16enjun w21enjun w26enjun;
w11enjun with w16enjun w21enjun w26enjun;
w16enjun with w21enjun w26enjun;
w21enjun with w26enjun;
! [univx1 univx6 univx11 univx16 univx21 univx26];
I S | univx1@0 univx6@1 univx11@2 univx16@3 univx21@4 univx26@5;
[I];
MODEL CONSTRAINT: L1 = 3 - L2 - L3;
T1 = 0 - T2 - T3;
output: sampstat residual stdyx tech4 modindices (5);
```

## Higher Order LGM: Curve of Factor Model









Number of Free Parameters		78						
Loglikelihood								
H0 Value		-10821.321						
H1 Value		-10686.827						
Information Criteria		S	WITH					
		I		-0.003	0.033	-0.077	0.939	
Akaike (AIC)		21798.643						
Bayesian (BIC)		22112.472						
Sample-Size Adjusted BIC		21864.960						
(n* = (n + 2) / 24)		Means						
		I		6.158	0.087	71.160	0.000	
		S		-0.053	0.018	-2.998	0.003	
Chi-Square Test of Model Fit		Variances						
Value		268.989		I	1.986	0.210	9.454	0.000
Degrees of Freedom		111		S	0.053	0.010	5.249	0.000
P-Value		0.0000						
RMSEA (Root Mean Square Error Of Approximation)				Residual Variances				
Estimate		0.059		UNIVX1	1.444	0.183	7.909	0.000
90 Percent C.I.		0.050 0.068		UNIVX6	0.878	0.120	7.297	0.000
Probability RMSEA <= .05		0.054		UNIVX11	1.171	0.138	8.512	0.000
CFI/TLI				UNIVX16	0.421	0.070	5.994	0.000
				UNIVX21	0.644	0.089	7.209	0.000
				UNIVX26	0.148	0.080	1.846	0.065
CFI		0.970						
TLI		0.959						
Chi-Square Test of Model Fit for the Baseline Model								
Value		5487.539						
Degrees of Freedom		153						
P-Value		0.0000						
SRMR (Standardized Root Mean Square Residual)								
Value		0.076						













































## LGM with Intercept and Slope Regressed on Gender (Time Invariant Covariate)



Mplus - [linear\_semesterdrinking8.inp]

File Edit View Mplus Graph Window Help


































































































Title: LGM Drinking three-month blocks - Paul Tremblay;  
 data: File is drinking\_semesters\_mpluspaper.dat;  
 Format is 1F4, 1F1, 9F8.2;  
 variable: names are studid gender drinker d1 d2 d3 d4 d5 d6 d7 d8;  
 missing = blank;  
 usevariables are d1 d2 d3 d4 d5 d6 d7 d8 gender\_d;  
 define:  
 if (gender eq 1) then gender\_d = 1;  
 if (gender eq 2) then gender\_d = 0;  
 analysis: estimator=mlr;  
 model: i s | d1@0 d2@0.143 d3@0.286 d4@0.429 d5@0.572 d6@0.715 d7@0.858 d8@1;  
 i s on gender\_d;  
 plot:  
 type=plot3;  
 series is d1 (0) d2 (1) d3 (2) d4 (3) d5 (4) d6 (5) d7 (6) d8 (7);  
 output: sampstat standardized tech1 tech3 tech4;

Ready

Mplus - [linear\_semesterdrinking8.out]

File Edit View Mplus Graph Window Help

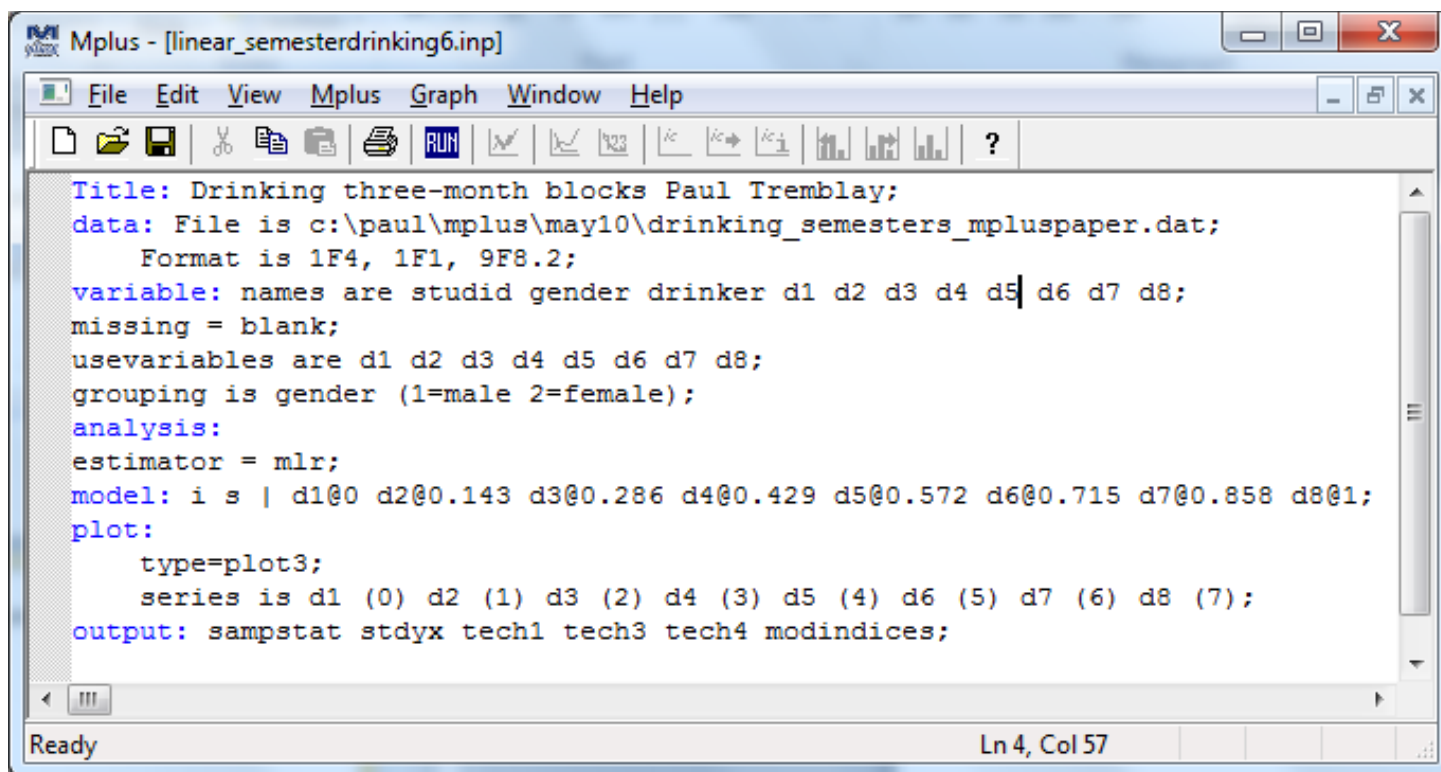
Mplus - [linear\_semesterdrinking8.out]

File Edit View Mplus Graph Window Help

I	ON				
	GENDER_D	3.884	0.870	4.462	0.000
S	ON				
	GENDER_D	-1.587	0.952	-1.667	0.095
S	WITH				
	I	-12.411	6.048	-2.052	0.040

Ready Ln 265, Col 31

## LGM Multiple groups (Gender)

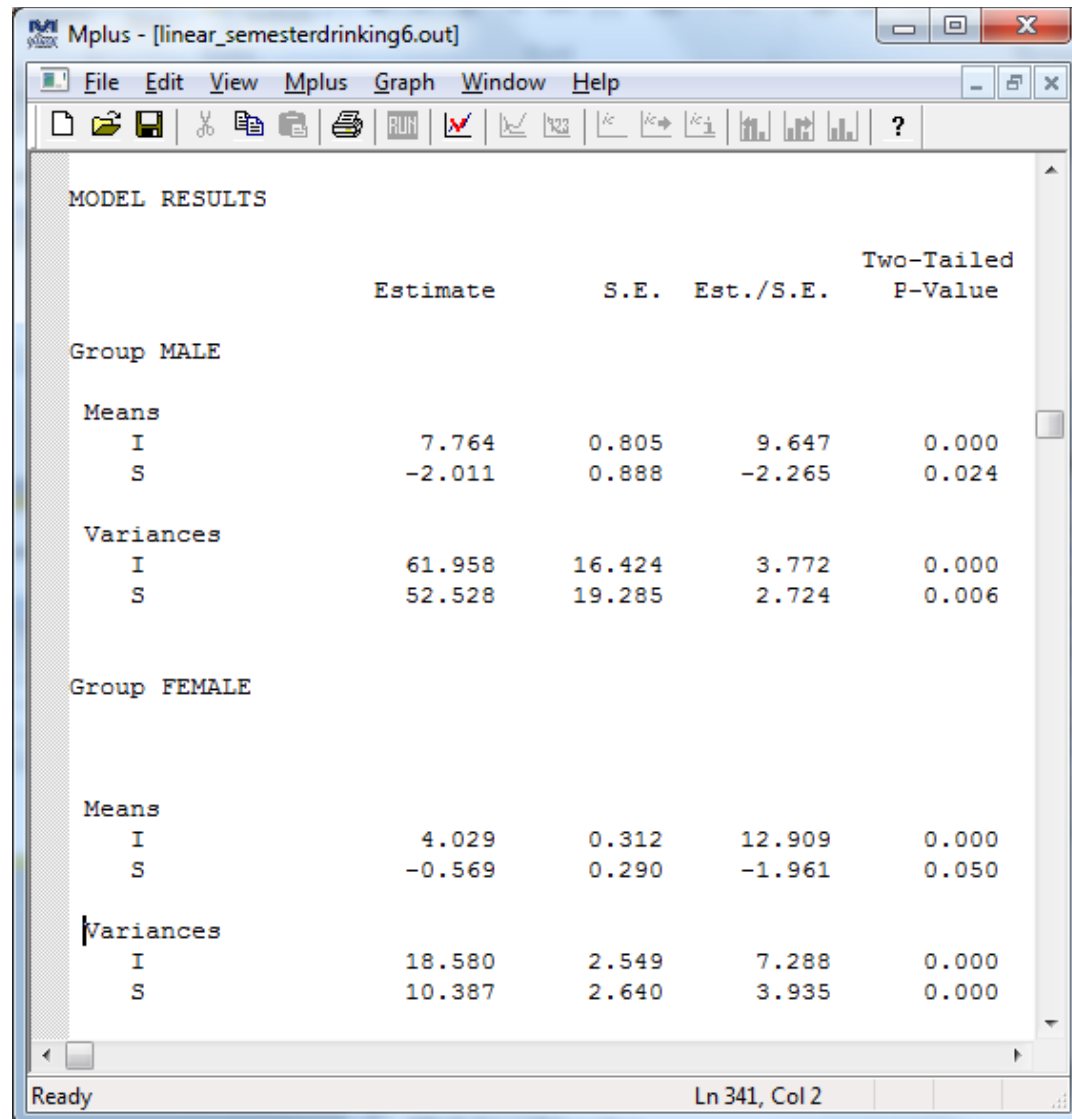


The screenshot shows the Mplus software window titled "Mplus - [linear\_semesterdrinking6.inp]". The menu bar includes File, Edit, View, Mplus, Graph, Window, and Help. The toolbar contains icons for file operations, running, and plotting. The main text area displays the following syntax:

```
Title: Drinking three-month blocks Paul Tremblay;
data: File is c:\paul\mplus\may10\drinking_semesters_mpluspaper.dat;
      Format is 1F4, 1F1, 9F8.2;
variable: names are studid gender drinker d1 d2 d3 d4 d5 d6 d7 d8;
missing = blank;
usevariables are d1 d2 d3 d4 d5 d6 d7 d8;
grouping is gender (1=male 2=female);
analysis:
  estimator = mlr;
model: i s | d1@0 d2@0.143 d3@0.286 d4@0.429 d5@0.572 d6@0.715 d7@0.858 d8@1;
plot:
  type=plot3;
  series is d1 (0) d2 (1) d3 (2) d4 (3) d5 (4) d6 (5) d7 (6) d8 (7);
output: sampstat stdyx tech1 tech3 tech4 modindices;
```

The status bar at the bottom indicates "Ready" and "Ln 4, Col 57".

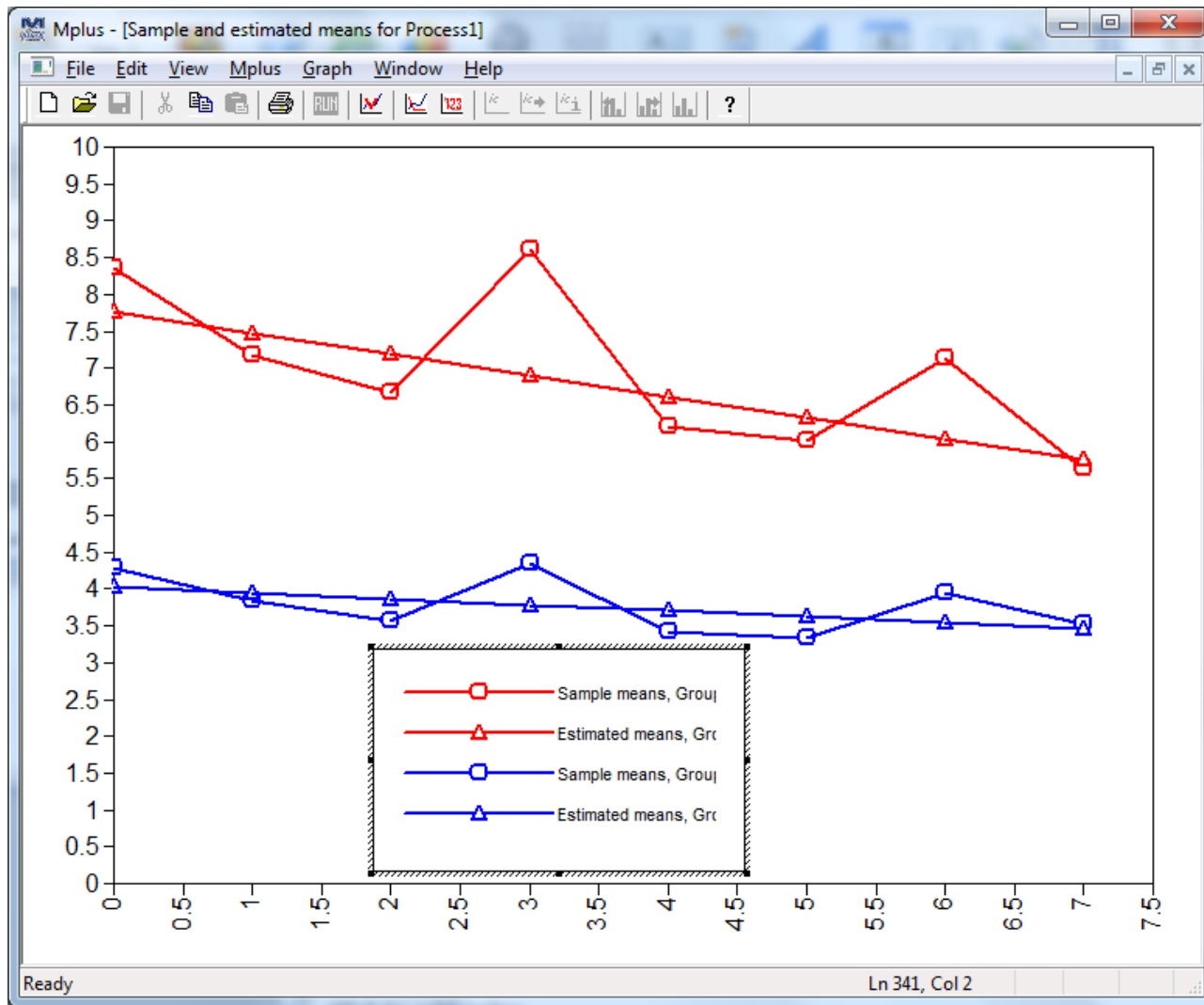
## LGM Multiple groups (Gender)



The screenshot shows the Mplus software window titled "Mplus - [linear\_semesterdrinking6.out]". The menu bar includes File, Edit, View, Mplus, Graph, Window, and Help. The toolbar contains icons for file operations, running, and plotting. The main window displays the "MODEL RESULTS" output, which is organized by group (MALE and FEMALE) and then by parameter type (Means and Variances). The output table has five columns: Estimate, S.E., Est./S.E., and Two-Tailed P-Value. The status bar at the bottom indicates "Ready" and "Ln 341, Col 2".

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Group MALE				
Means				
I	7.764	0.805	9.647	0.000
S	-2.011	0.888	-2.265	0.024
Variances				
I	61.958	16.424	3.772	0.000
S	52.528	19.285	2.724	0.006
Group FEMALE				
Means				
I	4.029	0.312	12.909	0.000
S	-0.569	0.290	-1.961	0.050
Variances				
I	18.580	2.549	7.288	0.000
S	10.387	2.640	3.935	0.000

## LGM Multiple groups (Gender)



## Measures within Persons

```
Title: PSY9555 Regression examples;
!note that two outliers of 0 were removed in average;
data: File is sem_plus2.dat;
      Format is 1F4, 1F1, 1F2, 23F8.2, 1F11.3, 72F8.2;
data widetolong: <
wide = drink1 drink2 drink3 drink4;
long = drink;
idvariable = person;
repetition = time;
variable: names are studid gender age
bppa bpv bpa bph bptot
sq1 sq2 sq3 sq4 sq5 sq6 sq7 sq8 sq9 sq10 sq11 sq12 sq13 sq14 sq15
es es_pt es_fin grade
drink1 drink2 drink3 drink4 epis1 epis2 epis3 epis4
stress1 stress2 stress3 stress4 pleased1 pleased2 pleased3 pleased4
enjoyc1 enjoyc2 enjoyc3 enjoyc4 enjoyu1 enjoyu2 enjoyu3 enjoyu4
effort1 effort2 effort3 effort4 harm1 harm2 harm3 harm4 dep1 dep2 dep3 dep4
drink1b drink2b drink3b drink4b epis1b epis2b epis3b epis4b
stress1b stress2b stress3b stress4b please1b please2b please3b please4b
enjoyc1b enjoyc2b enjoyc3b enjoyc4b enjoyu1b enjoyu2b enjoyu3b enjoyu4b
effort1b effort2b effort3b effort4b harm1b harm2b harm3b harm4b
dep1b dep2b dep3b dep4b;
missing = blank;
usevariables are drink person time;
cluster = person;
within = time;
analysis:
type = twolevel random;
model:
%within%
s | drink on time;
%between%
s with drink;
output: sampstat tech1;
```

When your data file is structured in the conventional one line per subject with repeated measures on the same line

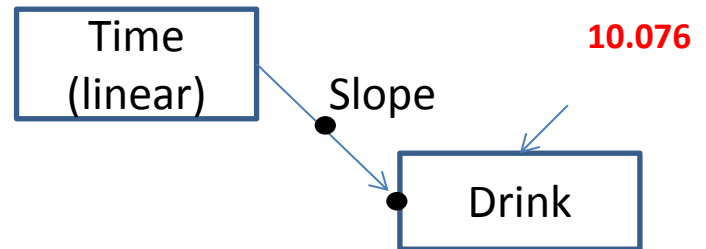
## Measures within Persons

### Loglikelihood

H0 Value -4756.126  
 H0 Scaling Correction Factor 3.5791  
 for MLR

### Information Criteria

Akaike (AIC) 9524.253  
 Bayesian (BIC) 9556.568  
 Sample-Size Adjusted BIC 9537.507  
 ( $n^* = (n + 2) / 24$ )



### MODEL RESULTS

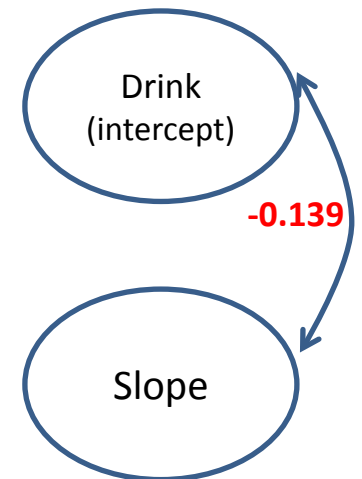
	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
<b>Within Level</b>				
Residual Variances				
DRINK	10.076	1.428	7.056	0.000
<b>Between Level</b>				
S WITH				
DRINK	-0.139	1.168	-0.119	0.905
<b>Means</b>				
DRINK	5.236	0.306	17.090	0.000
S	-0.118	0.085	-1.396	0.163
<b>Variances</b>				
DRINK	31.765	4.088	7.769	0.000
S	0.782	0.478	1.636	0.102

WITHIN

BETWEEN

Mean = 5.236  
 Var = 31.765

Mean = -.118  
 Var = .782





## Measures within Persons (Previous Model specified as LGM)

```
Title: PSY9555 Regression examples;
!note that two outliers of 0 were removed in average;
data: File is sem_mplus2.dat;
      Format is 1F4, 1F1, 1F2, 23F8.2, 1F11.3, 72F8.2;
!LISTWISE = ON;
variable: names are studid gender age
bppa bpv bpa bph bptot
sq1 sq2 sq3 sq4 sq5 sq6 sq7 sq8 sq9 sq10 sq11 sq12 sq13 sq14 sq15
es es_pt es_fin grade
drink1 drink2 drink3 drink4 epis1 epis2 epis3 epis4
stress1 stress2 stress3 stress4 pleased1 pleased2 pleased3 pleased4
enjoyc1 enjoyc2 enjoyc3 enjoyc4 enjoyu1 enjoyu2 enjoyu3 enjoyu4
effort1 effort2 effort3 effort4 harm1 harm2 harm3 harm4 dep1 dep2 dep3 dep4
drink1b drink2b drink3b drink4b epis1b epis2b epis3b epis4b
stress1b stress2b stress3b stress4b please1b please2b please3b please4b
enjoyc1b enjoyc2b enjoyc3b enjoyc4b enjoyu1b enjoyu2b enjoyu3b enjoyu4b
effort1b effort2b effort3b effort4b harm1b harm2b harm3b harm4b
dep1b dep2b dep3b dep4b;
missing = blank;
usevariables are drink1 drink2 drink3 drink4;
analysis:
type = general;
estimator = mlr;
model:
I S | drink1@0 drink2@1 drink3@2 drink4@3;
drink1-drink4 (1);
plot:
type is plot3;
series = drink1 (0) drink2 (1) drink3 (2) drink4 (3);
output: sampstat residual stdyx tech4 modindices;
```

Back to the conventional  
data structure specification

IN LGM these residuals are  
usually not constrained to  
equality but they are in  
MLM. I constrained them  
here.

## Example: Measures within Persons (LGM and MLM)

### LGM

Mplus - [lgm.out]

File Edit View Mplus Graph Window Help

MODEL RESULTS

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
S WITH				
I	-0.142	1.167	-0.122	0.903
Means				
I	5.238	0.307	17.088	0.000
S	-0.119	0.085	-1.404	0.160
Variances				
I	31.784	4.093	7.765	0.000
S	0.776	0.477	1.626	0.104
Residual Variances				
DRINK1	10.082	1.430	7.052	0.000
DRINK2	10.082	1.430	7.052	0.000
DRINK3	10.082	1.430	7.052	0.000
DRINK4	10.082	1.430	7.052	0.000

Ready Ln 219, Col 2

### MLM

Mplus - [lgm\_mlm.out]

File Edit View Mplus Graph Window Help

MODEL RESULTS

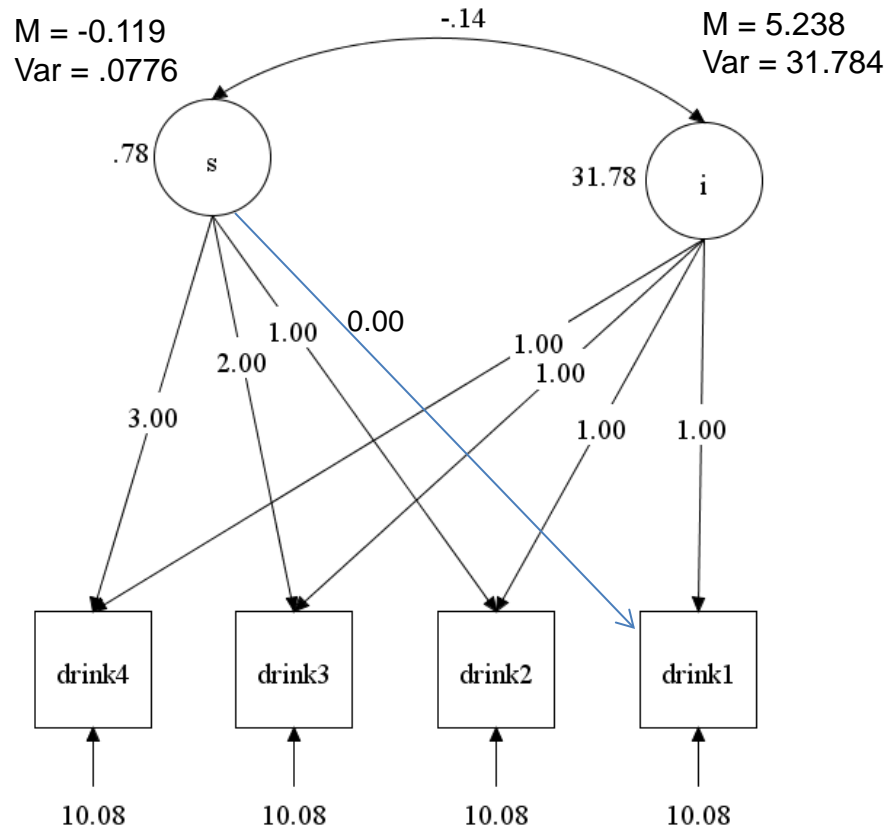
	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Within Level				
Residual Variances				
DRINK	10.076	1.428	7.056	0.000
Between Level				
S WITH				
DRINK	-0.139	1.168	-0.119	0.905
Means				
DRINK	5.236	0.306	17.090	0.000
S	-0.118	0.085	-1.396	0.163
Variances				
DRINK	31.765	4.088	7.769	0.000
S	0.782	0.478	1.636	0.102

Ready Ln 1, Col 1

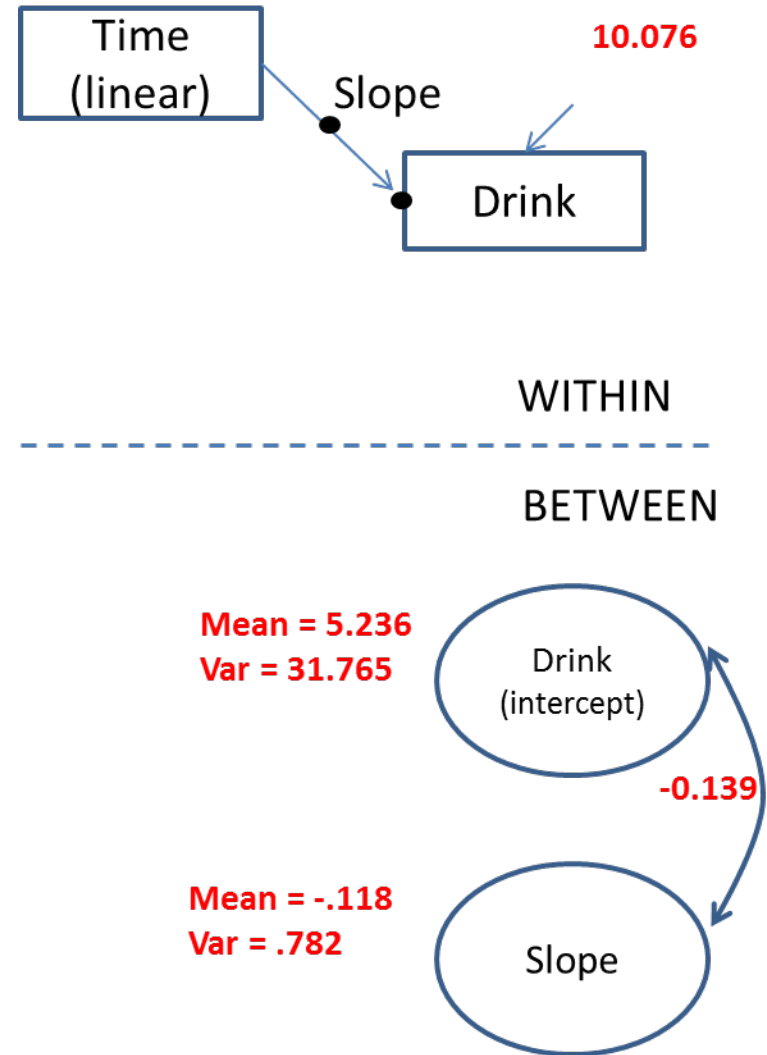
Same as in slide 13

## Example: Measures within Persons (LGM and MLM)

LGM



MLM



## Example: Measures within Persons (LGM and MLM)

### LGM

#### MODEL FIT INFORMATION

Number of Free Parameters 6

#### Loglikelihood

H0 Value	-4756.126
H0 Scaling Correction Factor for MLR	3.5794
H1 Value	-4731.180
H1 Scaling Correction Factor for MLR	3.6974

#### Information Criteria

Akaike (AIC)	9524.252
Bayesian (BIC)	9548.407
Sample-Size Adjusted BIC ( $n^* = (n + 2) / 24$ )	9529.368

#### Chi-Square Test of Model Fit

Value	13.178*
Degrees of Freedom	8
P-Value	0.1059
Scaling Correction Factor for MLR	3.7860

### MLM

#### MODEL FIT INFORMATION

Number of Free Parameters 6

#### Loglikelihood

H0 Value	-4756.126
H0 Scaling Correction Factor for MLR	3.5791

#### Information Criteria

Akaike (AIC)	9524.253
Bayesian (BIC)	9556.568
Sample-Size Adjusted BIC ( $n^* = (n + 2) / 24$ )	9537.507